18.440: Lecture 21

More continuous random variables

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Cauchy distribution

Cauchy distribution

Defining gamma function Γ

- ▶ Last time we found that if X is geometric with rate 1 and $n \ge 0$ then $E[X^n] = \int_0^\infty x^n e^{-x} dx = n!$.
- This expectation E[Xⁿ] is actually well defined whenever n > −1. Set α = n + 1. The following quantity is well defined for any α > 0: Γ(α) := E[X^{α−1}] = ∫₀[∞] x^{α−1}e^{-x}dx = (α − 1)!.
- So Γ(α) extends the function (α − 1)! (as defined for strictly positive integers α) to the positive reals.
- Vexing notational issue: why define Γ so that Γ(α) = (α − 1)! instead of Γ(α) = α!?
- At least it's kind of convenient that Γ is defined on (0,∞) instead of (−1,∞).

- ► The sum X of n independent geometric random variables of parameter p is negative binomial with parameter (n, p).
- Waiting for the *n*th heads. What is $P{X = k}$?
- Answer: $\binom{k-1}{n-1}p^{n-1}(1-p)^{k-n}p$.
- What's the continuous (Poisson point process) version of "waiting for the *n*th event"?

Poisson point process limit

- Recall that we can approximate a Poisson process of rate λ by tossing N coins per time unit and taking p = λ/N.
- Let's fix a rational number x and try to figure out the probability that that the nth coin toss happens at time x (i.e., on exactly xNth trials, assuming xN is an integer).
- Write $p = \lambda/N$ and k = xN. (Note $p = \lambda x/k$.)
- For large N, $\binom{k-1}{n-1}p^{n-1}(1-p)^{k-n}p$ is

$$\frac{(k-1)(k-2)\dots(k-n+1)}{(n-1)!}p^{n-1}(1-p)^{k-n}p$$

$$\approx \frac{k^{n-1}}{(n-1)!}p^{n-1}e^{-x\lambda}p = \frac{1}{N}\Big(\frac{(\lambda x)^{(n-1)}e^{-\lambda x}\lambda}{(n-1)!}\Big).$$

- ► The probability from previous side, $\frac{1}{N} \left(\frac{(\lambda x)^{(n-1)} e^{-\lambda x} \lambda}{(n-1)!} \right)$ suggests the form for a continuum random variable.
- Replace n (generally integer valued) with α (which we will eventually allow be to be any real number).
- Say that random variable X has gamma distribution with parameters (α, λ) if $f_X(x) = \begin{cases} \frac{(\lambda x)^{\alpha-1}e^{-\lambda x}\lambda}{\Gamma(\alpha)} & x \ge 0\\ 0 & x < 0 \end{cases}$.
- Waiting time interpretation makes sense only for integer α, but distribution is defined for general positive α.

Cauchy distribution

Cauchy distribution

- A standard Cauchy random variable is a random real number with probability density f(x) = ¹/_π ¹/_{1+x²}.
- There is a "spinning flashlight" interpretation. Put a flashlight at (0,1), spin it to a uniformly random angle in [-π/2, π/2], and consider point X where light beam hits the x-axis.

►
$$F_X(x) = P\{X \le x\} = P\{\tan \theta \le x\} = P\{\theta \le \tan^{-1}x\} = \frac{1}{2} + \frac{1}{\pi}\tan^{-1}x.$$

Find
$$f_X(x) = \frac{d}{dx}F(x) = \frac{1}{\pi}\frac{1}{1+x^2}$$
.

Cauchy distribution: Brownian motion interpretation

- The light beam travels in (randomly directed) straight line. There's a windier random path called Brownian motion.
- If you do a simple random walk on a grid and take the grid size to zero, then you get Brownian motion as a limit.
- We will not give a complete mathematical description of Brownian motion here, just one nice fact.
- ► FACT: start Brownian motion at point (x, y) in the upper half plane. Probability it hits negative x-axis before positive x-axis is ¹/₂ + ¹/_π tan⁻¹ ^y/_x. Linear function of angle between positive x-axis and line through (0,0) and (x, y).
- Start Brownian motion at (0,1) and let X be the location of the first point on the x-axis it hits. What's P{X < a}?</p>
- ► Applying FACT, translation invariance, reflection symmetry: $P\{X < x\} = P\{X > -x\} = \frac{1}{2} + \frac{1}{\pi} \tan^{-1} \frac{1}{x}.$
- ► So X is a standard Cauchy random variable.

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Question: what if we start at (0,2)?

- Start at (0, 2). Let Y be first point on x-axis hit by Brownian motion. Again, same probability distribution as point hit by flashlight trajectory.
- Flashlight point of view: Y has the same law as 2X where X is standard Cauchy.
- Brownian point of view: Y has same law as X₁ + X₂ where X₁ and X₂ are standard Cauchy.
- ▶ But wait a minute. Var(Y) = 4Var(X) and by independence Var(X₁ + X₂) = Var(X₁) + Var(X₂) = 2Var(X₂). Can this be right?
- Cauchy distribution doesn't have finite variance or mean.
- Some standard facts we'll learn later in the course (central limit theorem, law of large numbers) don't apply to it.

Cauchy distribution

Cauchy distribution

Beta distribution

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Beta distribution: Alice and Bob revisited

- Suppose I have a coin with a heads probability p that I don't know much about.
- What do I mean by not knowing anything? Let's say that I think p is equally likely to be any of the numbers {0, .1, .2, .3, .4, ..., .9, 1}.
- Now imagine a multi-stage experiment where I first choose p and then I toss n coins.
- ▶ Given that number h of heads is a 1, and b 1 tails, what's conditional probability p was a certain value x?

•
$$P\left(p=x|h=(a-1)\right) = \frac{\frac{1}{11}\binom{a}{a-1}x^{a-1}(1-x)^{b-1}}{P\{h=(a-1)\}}$$
 which is $x^{a-1}(1-x)^{b-1}$ times a constant that doesn't depend on x .

- Suppose I have a coin with a heads probability p that I really don't know anything about. Let's say p is uniform on [0, 1].
- ▶ Now imagine a multi-stage experiment where I first choose p uniformly from [0, 1] and then I toss n coins.
- ► If I get, say, a 1 heads and b 1 tails, then what is the conditional probability density for p?
- ► Turns out to be a constant (that doesn't depend on x) times $x^{a-1}(1-x)^{b-1}$.
- $\frac{1}{B(a,b)}x^{a-1}(1-x)^{b-1}$ on [0,1], where B(a,b) is constant chosen to make integral one. Can be shown that $B(a,b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$.
- ▶ What is *E*[*X*]?

• Answer:
$$\frac{a}{a+b}$$
.

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