### 18.175: Lecture 11

## Independent sums and large deviations

Scott Sheffield

MIT

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Large deviations

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- ▶ First Borel-Cantelli lemma: If  $\sum_{n=1}^{\infty} P(A_n) < \infty$  then  $P(A_n \text{ i.o.}) = 0.$
- ▶ Second Borel-Cantelli lemma: If  $A_n$  are independent, then  $\sum_{n=1}^{\infty} P(A_n) = \infty$  implies  $P(A_n \text{ i.o.}) = 1$ .

- ► Consider sequence of random variables  $X_n$  on some probability space. Write  $\mathcal{F}'_n = \sigma(X_n, X_{n_1}, ...)$  and  $\mathcal{T} = \bigcap_n \mathcal{F}'_n$ .
- $\mathcal{T}$  is called the **tail**  $\sigma$ -**algebra**. It contains the information you can observe by looking only at stuff arbitrarily far into the future. Intuitively, membership in tail event doesn't change when finitely many  $X_n$  are changed.
- Event that X<sub>n</sub> converge to a limit is example of a tail event. Other examples?
- ▶ **Theorem:** If  $X_1, X_2, ...$  are independent and  $A \in \mathcal{T}$  then  $P(A) \in \{0, 1\}.$

▶ **Thoerem:** Suppose  $X_i$  are independent with mean zero and finite variances, and  $S_n = \sum_{i=1}^n X_n$ . Then

$$P(\max_{1\leq k\leq n}|S_k|\geq x)\leq x^{-2}\mathrm{Var}(S_n)=x^{-2}E|S_n|^2.$$

Main idea of proof: Consider first time maximum is exceeded. Bound below the expected square sum on that event.

### Kolmogorov three-series theorem

- **Theorem:** Let  $X_1, X_2, \ldots$  be independent and fix A > 0. Write  $Y_i = X_i \mathbb{1}_{\{|X_i| \le A\}}$ . Then  $\sum X_i$  converges a.s. if and only if the following are all true:
  - ►  $\sum_{n=1}^{\infty} P(|X_n| > A) < \infty$ ►  $\sum_{n=1}^{\infty} EY_n$  converges ►  $\sum_{n=1}^{\infty} Var(Y_n) < \infty$
- Main ideas behind the proof: Kolmogorov zero-one law implies that  $\sum X_i$  converges with probability  $p \in \{0, 1\}$ . We just have to show that p = 1 when all hypotheses are satisfied (sufficiency of conditions) and p = 0 if any one of them fails (necessity).
- To prove sufficiency, apply Borel-Cantelli to see that probability that  $X_n \neq Y_n$  i.o. is zero. Subtract means from  $Y_n$ , reduce to case that each  $Y_n$  has mean zero. Apply Kolmogorov maximal inequality.

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## Recall: moment generating functions

- Let X be a random variable.
- ► The moment generating function of X is defined by M(t) = M<sub>X</sub>(t) := E[e<sup>tX</sup>].
- When X is discrete, can write M(t) = ∑<sub>x</sub> e<sup>tx</sup> p<sub>X</sub>(x). So M(t) is a weighted average of countably many exponential functions.
- When X is continuous, can write M(t) = ∫<sup>∞</sup><sub>-∞</sub> e<sup>tx</sup> f(x)dx. So M(t) is a weighted average of a continuum of exponential functions.
- We always have M(0) = 1.
- If b > 0 and t > 0 then  $E[e^{tX}] \ge E[e^{t\min\{X,b\}}] \ge P\{X \ge b\}e^{tb}.$
- If X takes both positive and negative values with positive probability then M(t) grows at least exponentially fast in |t| as |t| → ∞.

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# Recall: moment generating functions for i.i.d. sums

- We showed that if Z = X + Y and X and Y are independent, then  $M_Z(t) = M_X(t)M_Y(t)$
- If X<sub>1</sub>...X<sub>n</sub> are i.i.d. copies of X and Z = X<sub>1</sub> + ... + X<sub>n</sub> then what is M<sub>Z</sub>?
- Answer:  $M_X^n$ . Follows by repeatedly applying formula above.
- This a big reason for studying moment generating functions. It helps us understand what happens when we sum up a lot of independent copies of the same random variable.

- Consider i.i.d. random variables  $X_i$ . Want to show that if  $\phi(\theta) := M_{X_i}(\theta) = E \exp(\theta X_i)$  is less than infinity for some  $\theta > 0$ , then  $P(S_n \ge na) \to 0$  exponentially fast when  $a > E[X_i]$ .
- ► Kind of a quantitative form of the weak law of large numbers. The empirical average A<sub>n</sub> is very unlikely to ε away from its expected value (where "very" means with probability less than some exponentially decaying function of n).
- Write γ(a) = lim<sub>n→∞</sub> <sup>1</sup>/<sub>n</sub> log P(S<sub>n</sub> ≥ na). It gives the "rate" of exponential decay as a function of a.

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