# 6.241 Dynamic Systems and Control

Lecture 2: Least Square Estimation

Readings: DDV, Chapter 2

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# Outline

Least Squares Estimation

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# Least Squares Estimation

• Consider an system of m equations in n unknown, with m > n, of the form

$$y = Ax$$
.

- Assume that the system is inconsistent: there are more equations than unknowns, and these equations are non linear combinations of one another.
- In these conditions, there is no x such that y Ax = 0. However, one can write e = y Ax, and find x that minimizes ||e||.
- In particular, the problem

$$\min_{x} \|e\|_2 = \min_{x} \|y - Ax\|_2$$

is a least squares problem. The optimal x is the least squares estimate.

# Computing the Least-Square Estimate

- The set M := {z ∈ ℝ<sup>m</sup> : z = Ax, x ∈ ℝ<sup>n</sup>} is a subspace of ℝ<sup>m</sup>, called the range of A, R(A), i.e., the set of all vectors that can be obtained by linear combinations of the columns of A.
- Recall the projection theorem. Now we are looking for the element of *M* that is "closest" to *y*, in terms of 2-norm. We know the solution is such that

$$\mathsf{e} = (\mathsf{y} - A\hat{x}) \perp \mathcal{R}(A).$$

• In particular, if  $a_i$  is the *i*-th column of A, it is also the case that

$$(y - A\hat{x}) \perp \mathcal{R}(A) \quad \Leftrightarrow \quad a'_i(y - A\hat{x}) = 0, \qquad i = 1, \dots, n$$
  
 $A'(y - A\hat{x}) = 0$   
 $A'A\hat{x} = A'y$ 

• A'A is a  $n \times n$  matrix; is it invertible? It if were, then at this point it is easy to recover the least-square solution as

$$\hat{x} = (A'A)^{-1}A'y.$$

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# The Gram product

- Let us take a more abstract look at this problem, e.g., to address the case that the data vector y is infinite-dimensional.
- Given an array of  $n_A$  vectors  $A = [a_1| \dots |a_{n_A}]$ , and an array of  $n_B$  vectors  $B = [b_1| \dots |b_{n_B}]$ , both from an inner vector space V, define the Gram Product  $\prec A, B \succ$  as a  $n_A \times n_B$  matrix such that its (i, j) entry is  $\langle a_i, b_i \rangle$ .
- For the usual Euclidean inner product in an *m*-dimensional space,

$$\prec A, B \succ = A'B.$$

• Symmetry and linearity of the inner product imply symmetry and linearity of the Gram product.

### The Least Squares Estimation Problem

• Consider again the problem of computing

$$\min_{\mathbf{x}\in\mathbb{R}^n} \|\underbrace{\mathbf{y}-\mathbf{A}\mathbf{x}}_{e}\| = \min_{\hat{\mathbf{y}}\in\mathcal{R}(\mathbf{A})} \|\mathbf{y}-\hat{\mathbf{y}}\|.$$

- y can be an infinite-dimensional vector—as long as n is finite.
- We assume that the columns of  $A = [a_1, a_2, \dots, a_n]$  are independent.

#### Lemma (Gram matrix)

The columns of a matrix A are independent  $\Leftrightarrow \prec A, A \succ$  is invertible.

**Proof**— If the columns are dependent, then there is  $\eta \neq 0$  such that  $A\eta = \sum_j a_j \eta_j = 0$ . But then  $\sum_j \langle a_i, a_j \rangle \eta j = 0$  by the linearity of inner product. That is,  $\prec A, A \succ \eta = 0$ , and hence  $\prec A, A \succ$  is not invertible. Conversely, if  $\prec A, A \succ \eta = 0$ , and hence  $\prec A, A \succ \eta = 0$  for some  $\eta \neq 0$ . In other words  $\eta' \prec A, A \succ \eta = 0$ , and hence  $A\eta = 0$ .

### The Projection theorem and least squares estimation 1

- y has a unique decomposition  $y = y_1 + y_2$ , where  $y_1 \in \mathcal{R}(A)$ , and  $y_2 \in \mathcal{R}^{\perp}(A)$ .
- To find this decomposition, let  $y_1 = A\alpha$ , for some  $\alpha \in \mathbb{R}^n$ . Then, ensure that  $y_2 = y y_1 \in \mathcal{R}^{\perp}(A)$ . For this to be true,

$$\langle a_i, y - A\alpha \rangle = 0, \quad i = 1, \dots, n,$$

i.e.,

$$\prec A, y - A\alpha \succ = 0.$$

• Rearranging, we get

$$\prec A, A \succ \alpha = \prec A, y \succ$$

• if the columns of A are independent,

$$\alpha = \prec A, A \succ^{-1} \prec A, y \succ$$

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### The Projection theorem and least squares estimation 2

- Decompose  $e = e_1 + e_2$  similarly  $(e_1 \in \mathcal{R}(A)$ , and  $e_2 \in \mathcal{R}^{\perp}(A))$ .
- Note  $||e||^2 = ||e_1||^2 + ||e_2||^2$ .
- Rewrite e = y Ax as

$$e_1 + e_2 = y_1 + y_2 - Ax,$$

i.e.,

$$e_2 - y_2 = y_1 - e_1 - Ax.$$

- Each side must be 0, since they are on orthogonal subspaces!
- $e_2 = y_2$ —can't do anything about it.

•  $e_1 = y_1 - Ax = A(\alpha - x)$ —minimize by choosing  $x = \alpha$ . In other words

$$\hat{x} = \prec A, A \succ^{-1} \prec A, y \succ .$$

# Examples

• If  $y, e \in \mathbb{R}^m$ , and it is desired to minimize  $||e||^2 = e'e = \sum_{i=1}^m |e_i|^2$ , then  $\hat{x} = (A'A)^{-1}A'y$ 

(If the columns of A are mutually orthogonal, A'A is diagonal, and inversion is easy)

• if  $y, e \in \mathbb{R}^m$ , and it is desired to minimize e'Se, where S is a Hermitian, positive-definite matrix, then

$$\hat{x} = (A'SA)^{-1}A'Sy.$$

- Note that if S is diagonal, then  $e'Se = \sum_{i=1}^{m} s_{ii}|e_i|^2$ , i.e., we are minimizing a weighted least square criterion. A large  $s_{ii}$  penalizes the *i*-th component of the error more relative to the others.
- In a general stochastic setting, the weight matrix S should be related to the noise covariance, i.e.,

$$S = (E[ee'])^{-1}$$

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