1.010 Uncertainty in Engineering Fall 2008

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### **Application Example 17**

# Sums of iid Random Variables

#### **Central Limit Theorem**

The Central Limit Theorem establishes that, when properly normalized, the sum of *n* independent and identically distributed (iid) random variables  $X_i$  (i = 1, 2, ..., n) with finite variance approaches the standard normal distribution as  $n \to \infty$ . The standard normal distribution has zero mean value and unit variance and probability density function (PDF) given by:

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, \quad -\infty < z < \infty$$
 (1)

### The Case with Finite *n*

In practice, n is finite and the question arises whether n is large enough to assume that the sum has normal distribution. The rate at which the probability density function of the normalized sum

$$Z_n = \frac{\left(\sum_{i=1}^n X_i\right) - nm_X}{\sqrt{n\sigma_X}}$$
(2)

approaches the function in equation (1) depends on the distribution of X and the region of the distribution one considers. Convergence is usually fast near the mean value but may be slow in the tail regions, especially when the distribution of X is highly skewed.

To illustrate convergence, we consider the cases when X has uniform or exponential distribution.

### **1. Uniform Distribution**

Suppose that *X* has uniform distribution in the interval [-0.5, 0.5]. In this case the mean value and variance of *X* are  $m_X = 0$  and  $(\sigma_X)^2 = 1/12$ , and the normalized sum  $Z_n$  in equation (2) is:

$$Z_n = \sqrt{12} \frac{\left(\sum_{i=1}^n X_i\right)}{\sqrt{n}}$$
(3)

Since 
$$\sum_{i=1}^{n} X_i$$
 ranges from  $-0.5n$  to  $0.5n$ , the PDF of  $Z_n$  is 0 outside the interval  $[-\sqrt{3n}, \sqrt{3n}]$ 

For X uniformly distributed, the distribution of  $Z_n$  in equation (3) does not have known analytical form but can be obtained numerically, as follows. Consider first the case n = 2. The PDF of  $Y_2 = X_1 + X_2$  is given by

$$f_{Y_2}(y) = \int_{-\infty}^{\infty} f_X(x) f_X(y-x) dx$$
(4)

Integrals of the type in equation (4) are called convolution integrals. A convenient way to evaluate such integrals is to work with the characteristic function of *X*,  $\varphi_X(t)$ , which is defined as the Fourier transform of  $f_X$ :

$$\varphi_X(t) = \int_{-\infty}^{\infty} f_X(x) e^{-ixt} dx$$
(5)

with inverse transform

$$f_X(x) = \frac{1}{2\pi} \int_{-\infty}^{\infty} \varphi_X(t) e^{ixt} dt$$
(6)

One can show that the characteristic function of  $Y_2$  in equation (4) is simply the square of the characteristic function of X. More in general, the characteristic function of  $Y_n = \sum_{i=1}^n X_i$  is the  $n^{\text{th}}$ 

power of the characteristic function of *X*:

$$\varphi_{Y_n}(t) = \left[\varphi_X(t)\right]^n \tag{7}$$

Hence, a simple way to obtain the distribution of  $Z_n$  is to: (a) find  $\varphi_X(t)$ , either analytically or numerically, (b) raise  $\varphi_X(t)$  to the  $n^{\text{th}}$  power, (c) take the inverse Fourier transform to obtain  $f_{Y_n}$ , and (d) find  $f_{Z_n}(z)$  through:

$$f_{Z_n}(z) = \sqrt{n\sigma_X} f_{Y_n}(\sqrt{n\sigma_X} z + nm_X)$$
(8)

We have used this method to obtain the PDF of  $Z_n$  in equation (3) for different *n*. Figure 1 compares  $f_{Z_n}$  with the standard normal PDF in equation (1), for n = 2, 5, 10, and 20. For each value of *n*, the results for  $f_{Z_n}$  are shown in both arithmetic scale (left column) and log scale (right column). From the arithmetic plots, it appears that convergence to the normal density is achieved already for n = 5 and definitely for n = 10. However, this is true only up to about 3 or 4 standard deviations away from the mean. As the log plots show, beyond this central region the distributions of *Z* and  $Z_n$  are still different and one needs larger values of *n* for convergence. Notice that these extreme tail regions are associated with very small values of the PDF, and hence very small exceedance probabilities.

### 2. Exponential Distribution

Suppose now that X has exponential distribution with PDF

$$f_X(x) = e^{-x}, \ x \ge 0$$
 (9)

In this case  $m_X = (\sigma_X)^2 = 1$  and, using equation (2),  $Z_n$  is given by

$$Z_n = \frac{\left(\sum_{i=1}^n X_i\right) - n}{\sqrt{n}} \tag{10}$$

The PDF of  $Z_n$  can be found analytically, as follows. We know that the sum  $Y_n = \sum_{i=1}^n X_i$  has

gamma distribution with density

$$f_{Y_n}(y) = \frac{y^{n-1} e^{-y}}{(n-1)!}, \ y \ge 0$$
(11)

Then, using results for distributions of linear functions, the PDF of  $Z_n$  in equation (10) is

$$f_{Z_n}(z) = \frac{\sqrt{n}}{(n-1)!} \left( n + \sqrt{n}z \right)^{n-1} e^{-(n+\sqrt{n}z)}, \quad z \ge -\sqrt{n}$$
(12)

Figure 2 compares the density in equation (12) with the density in equation (1) for n = 2, 10, 20, 130. Convergence to the normal distribution is slower than in the case of the uniform

distribution. The reason is that the exponential distribution is highly skewed and one needs a large n to "undo" that skewness. This is especially clear from the log plots.

## Problem 17.1

- (a) Notice that, if X has uniform distribution,  $Z_2$  has symmetrical triangular distribution. Using the results in Figure 1, discuss how large n should be for convergence to the normal distribution in the range [-3, 3] if X has symmetrical triangular distribution.
- (b) Computers usually simulate normally distributed variables as the sum of 12 uniformly distributed variables. See what equation (3) becomes for n = 12. Suppose now that you need to accurately reproduce the tail of the normal distribution up to about 5 standard deviations from the mean. Using Figure 1, estimate a reasonable value of n that would give you the desired accuracy. How could you use the normal variable simulator of the computer (which uses n = 12) to simulate normal variables that meet your higher-accuracy criterion?



Figure 1: Comparison of the distributions of  $Z_n$  (solid lines) and Z (dashed lines) for different n, for the case when X has uniform distribution. The plots on the left are in arithmetic scale, whereas the plots on the right are in log scale and emphasize the tail regions of the distributions.



Figure 2: Comparison of the distributions of  $Z_n$  (solid lines) and Z (dashed lines) for different n, for the case when X has exponential distribution. The plots on the left are in arithmetic scale, whereas the plots on the right are in log scale and emphasize the tail regions of the distributions.